

ORIGINAL ARTICLE

Open Access

Changes in labour market transitions in Ireland over the Great Recession: what role for policy?

Adele Bergin^{1,2}, Elish Kelly^{1,2} and Seamus McGuinness^{1,2,3*}

* Correspondence: Seamus.
McGuinness@esri.ie

¹The Economic and Social Research
Institute, Whitaker Square, Sir John
Rogerson's Quay, Dublin 2, Ireland

²The Department of Economics,
Trinity College Dublin, Dublin,
Ireland

Full list of author information is
available at the end of the article

Abstract

This paper assesses the impact the Great Recession had on individuals' transitions to and from unemployment in Ireland. The rate of transition from unemployment to employment declined between 2006 and 2011, while the rate from employment to unemployment increased. The results indicate that young people are much less likely to exit unemployment but, at the same time, have a lower risk of becoming unemployed. Education has become an increasingly important factor in supporting unemployment exits and reducing the risk of unemployment since the recession. The findings are supportive of current policies aimed at facilitating access to higher education for unemployed people, but they raise questions with regard to the adequacy of the Youth Guarantee.

JEL classification: I21; J01; J60; J64; J68

Keywords: Labour market transitions; Great recession; Longitudinal data; Decomposition techniques; Ireland

1 Introduction

While few countries escaped the adverse impact of the Great Recession, Ireland experienced one of the most severe downturns in the industrialised world. Between 2008 and 2009, the country suffered a cumulative fall in Gross National Product of close to 9 per cent. By 2012 (Q3), real GNP was 8.5 per cent below its 2008 (Q1) level, which is a stark depiction of the rate of contraction experienced within the economy. This collapse in economic activity has had severe knock-on effects on Ireland's labour market. In particular, the unemployment rate increased from 4.6 per cent in 2004 to 15.0 per cent in 2012, while, as one would expect, the employment rate declined – falling from 65.9 per cent in 2004 to 58.8 per cent in 2012 (see Table 1).

The collapse in the property sector in Ireland, which was one of the main factors underlining its economic demise, resulted in most job losses occurring in the construction sector, and in particular among males. McGinnity *et al.* (2014) report that the number of people employed in the construction sector fell by just over 61 per cent between 2007 and 2012 – falling from 266,174 to 103,212, with the sector's share in total employment falling from 12.6 per cent in 2007 to 5.5 per cent in 2012. Apart from Information and Communication, Education, Health and the Arts, all other sectors experienced job losses, but none to the extent of the construction sector.

Table 1 Irish labour force statistics: 2004 to 2014

	2004 (Q2)	2005 (Q2)	2006 (Q2)	2007 (Q2)	2008 (Q2)	2009 (Q2)	2010 (Q2)	2011 (Q2)	2012 (Q2)	2013 (Q2)	2014 (Q1)
Unemployment rate %	4.6	4.7	4.6	4.8	5.7	12.2	13.9	14.6	15.0	13.9	12.0
Participation rate %	60.5	62.1	63.2	64.1	63.7	62.5	61.1	60.5	60.1	60.5	59.7
Employment rate % (persons aged 15–64)	65.9	67.5	68.5	69.1	67.9	62.2	60.0	59.2	58.8	60.2	60.8

Source: Quarterly National Household Survey, Central Statistics Office.

As the downturn progressed, the profile of unemployed people changed, particularly in terms of their educational composition. In 2006, prior to the recession, over two-thirds of those who were unemployed had a second-level qualification or less, with only a fifth having completed third-level. However, over the course of the recession, there has been a striking growth in the number of unemployed individuals with a Post Leaving Cert¹ (PLC) level qualification (which includes apprenticeships), while the proportion of unemployed females with a graduate qualification increased significantly from 22.5 per cent in 2006 to 30.8 per cent in 2011 (Kelly *et al.* 2012).

Existing literature on the impact of the recession on the Irish labour market has found that it has had a differential impact across age groups, with young people in particular being severely affected. Kelly *et al.* (2014) report that total employment among those aged 25 and less fell by almost 60 per cent, from 357,000 in Q3 2007 to 148,000 in Q1 2013. There has been a small increase in the number of employed youths since this time-point, with the figure standing at 160,500 in Q3 2014. The corollary of this fall in employment among young people has been a rapid rise in their unemployment rate, which has grown from 9 per cent in Q3 2007 to almost 30 per cent in Q3 2011, and also a rise in inactivity. Conefrey (2011) found that the majority of young people who exited the labour force after the recession in 2008, and who remained in Ireland, returned to education. However, Kelly *et al.* (2014) found that the rate of transition to employment of those who became unemployed fell considerably, from 38 per cent in 2006 to 17.4 per cent in 2011. Kelly *et al.* (2014) results showed that this fall in the transition rate was not due to changes in the composition of unemployed youths but to adjustments in the external economic environment that resulted in changes in the impact of possessing certain individual characteristics. For example, there was a rise in the marginal impact of education and Irish nationality on the probability of a successful transition from unemployment to employment. The number of young people not in employment, education or training, i.e., NEETs, has also grown since the recession, increasing from 11.8 per cent in 2006 to 24 per cent in 2011 (Eurostat, 2013).² Similar to unemployed youths, Kelly and McGuinness (2014) found that NEET individuals' transition rate to employment fell over the course of the recession as well – declining from 21.4 per cent in 2006 to 13.6 per cent in 2011.

Research by Barrett and Kelly (2012) shows that the recession has had a differential nationality effect and, in particular, that immigrants have suffered more than natives. They report that the annual rate of job loss was close to 20 per cent for immigrants in 2009 compared to about 7 per cent for Irish-born. In addition, Barrett and Kelly (2012) found that the employment probabilities of immigrants from the accession states that joined the European Union in May 2004 were particularly negatively affected, especially males from these countries.

In the context of the dramatic changes that have taken place in the Irish labour market since the Great Recession, this paper focuses on identifying the impact that various socio-economic characteristics (e.g., age, gender, marital status, etc.) and job traits (e.g., employment sector, trade union membership, etc.) have on the relative risk of a person losing their job (i.e., transitioning from employment to unemployment) or finding employment (i.e., transitioning from unemployment to employment) and, in particular, assesses how the impact of such characteristics have changed pre and post the Great Recession.³ In addition to examining the impact of various socio-economic characteristics on individuals' labour market transition patterns, we are also interested in investigating what has driven the change in the unemployment-to-employment and employment-to-unemployment transition rates between 2006 and 2011. This analysis is carried out using decomposition techniques.

The remainder of the paper is structured as follows. Section 2 describes the data that is used in the study. Descriptive statistics are presented in Section 3. The methodology and results are outlined in Sections 4, policy implications are discussed in Section 5, and Section 6 presents the main conclusions.

2 Data

Newly available longitudinal data from the *Quarterly National Household Survey* (QNHS), which is Ireland's Labour Force Survey,⁴ is used to conduct the analyses undertaken in this paper. The objective of the QNHS is to provide quarterly labour force estimates of employment, unemployment, etc. The survey is continuous and targets all private households. Each quarter, approximately 26,000 households are sampled.⁵ Households are asked to take part in the survey for five consecutive quarters. In each quarter, one-fifth of the households surveyed are replaced, and the QNHS sample involves an overlap of 80 per cent between consecutive quarters and 20 per cent between the same quarters in consecutive years. While participation in the QNHS is voluntary, the response rate is high (approximately 85 per cent).⁶

One of the main benefits of using the QNHS longitudinal data is that it allowed us to follow the labour market transitions of individuals for up to 5 consecutive quarters. In terms of the periods covered in our study, we focused on comparing individuals' labour market transitions both pre and post the Great Recession. Quarter 2 (Q2) 2006 was selected as the starting point for our pre-recession analysis, which was during the height of the boom in Ireland; and Q2 2011 marked the starting point for the analysis relating to the post-recession period.⁷ We selected two unbalanced panels and focused on working age individuals (age 15–64) who remained in the panel for, at least, the consecutive quarter. Within each panel, by following the same individuals over time, we can observe changes in labour market status and determine gross flows between different labour market states. The transition rate for moving into employment (unemployment) is calculated as the percentage of those who were employed (unemployed) in the previous quarter whose labour market status has changed to unemployed (employed) in the current quarter. Given that we only consider individuals who remain in Ireland over the observation periods, those who migrated from Ireland during the time period covered by each panel are excluded from the analysis.

In this paper, we begin by assessing the impact of various socio-economic and employment characteristics on an unemployed person's likelihood of transitioning to employment. We then examine the impact of characteristics on an employed person's

likelihood of becoming unemployed. Models are estimated for both 2006 and 2011. The characteristics investigated are as follows: gender, age, nationality, marital status, educational attainment, geographic location, sector of employment, previous unemployment duration, trade union membership, employment contract type, job type (i.e., full-time or part-time), firm size and job search methods.

3 Descriptive statistics

Table 2 shows average transition rates over the five waves for 2006 and 2011. In 2006, the pre-crisis period, over 37 per cent of those who were unemployed in the previous quarter transitioned into employment.⁸ In 2011, the average unemployment to employment

Table 2 Transition rates

	2006	2011
Overall:		
Employment to unemployment	1.8%	4.9%
Unemployment to employment	37.3%	27.8%
By age group:		
Employment to unemployment:		
15-19	3.3%	6.3%
20-24	3.4%	9.7%
25-34	2.0%	6.1%
35-44	1.4%	4.3%
45-54	1.2%	3.6%
55 plus	0.8%	2.6%
Unemployment to employment:		
15-19	19.5%	8.5%
20-24	43.4%	23.1%
25-34	43.3%	30.8%
35-44	38.7%	31.4%
45-54	38.8%	31.4%
55 plus	33.5%	26.9%
By education level:		
Employment to unemployment:		
Below leaving certificate	2.1%	6.0%
Leaving cert	1.9%	5.7%
PLC	1.6%	6.8%
Third level non-degree	1.5%	3.6%
Third level degree	1.2%	2.9%
Unemployment to employment:		
Below leaving certificate	24.3%	15.7%
Leaving cert	43.0%	27.5%
PLC	46.9%	27.2%
Third level non-degree	49.1%	40.7%
Third level degree	59.8%	50.0%

Note: Derived from the QNHS micro-data. The base for the employed to unemployed transition rates is employed in the previous quarter, and the base for the unemployed to employed transition rates is unemployed in the previous quarter.

transition rate had fallen by almost 10 percentage points. In 2006, when the economy was essentially at full employment, the transition rate from employment to unemployment was under 2 per cent, and this rate rose to just under 5 per cent in 2011.

There is an almost linear negative relationship between age and transitioning to unemployment. Apart from the youngest age group (15–19), the employment to unemployment transition rates approximately trebled for each age group in 2011 compared to 2006.

Younger workers have the lowest transition rates from unemployment to employment. The employment transition rates fell for all age groups in 2011, with the largest percentage point falls occurring for those under the age of 35.

Table 2 reveals that those with higher levels of education are less likely to move to unemployment, and this effect is more marked in 2011. There is a strong positive relationship between education level and transitioning into employment. In 2006, 60 per cent of unemployed graduates moved into employment while the comparable figure for those with less than Leaving Certificate education was only 24 per cent. In 2011, the education gradient persisted although the transition rates had fallen somewhat for each education level.

4 Methodology and results

With respect to assessing the link between measures of labour market status in period t and $t + 1$, we employ a random effects logit model.⁹ We estimate separate models for two distinct labour market transitions between the two periods: i) transitions from unemployment to employment and ii) transitions from employment to unemployment, with the equations taking the following general form:

$$Y_{it+1} = X'_{it}\beta + \alpha_i + u_{it} \quad (i = 1, \dots, N; t = 1, \dots, T), \quad (1)$$

where Y_{it+1} captures a person's labour market status between the previous quarter and the current quarter, X'_{it} is a vector of personal and job characteristics; α_i are the individual random effects; and u_{it} is an error term, which is assumed to be identically and independently distributed. The models are estimated on two unbalanced panels covering Q2 2006 to Q2 2007 and Q2 2011 to Q2 2012.¹⁰

The main weakness with the random effects model proposed here is that it assumes the explanatory variables are uncorrelated with the error term, suggesting that a fixed effects model, where individual-specific effects are captured through the inclusion of a separate intercept for each individual in the data, might be more appropriate. A conditional fixed effects logit model might have been estimated, but it comes at a huge cost: ii) the loss of many cases, since these models are conditional on changes in the outcome variable being observed, which is a particular problem when using quarterly panel data; and (ii) since most of the explanatory variables will not change over five quarters, most of our controls relating to, for example, education, region and industrial sector will drop out of the models. For these reasons, we are forced to adopt the random effects logit model. Nevertheless, as we are interested in changes in the value of observable characteristics over time, and accepting that any bias relating to unobserved heterogeneity is likely to affect both panels in similar ways, we are less concerned with respect to potential correlations with the error term in our current context.

Table 3 shows the results from models estimating the likelihood of transitioning from unemployment to employment in both 2006 and 2011. The odds ratios for each observable characteristic are reported and a number of notable changes in the importance of certain attributes are apparent pre and post recession. In 2006, relative to the comparison case of individual's aged 15 to 19, older workers had, on average, 4 to 5 times the odds of transitioning into employment, with the effect being relatively constant throughout the age ranges. By 2011, relative to the base case, workers aged 20 to 24 had almost ten times the odds of successfully exiting to employment, while individuals in the older age brackets had between 27 and 45 times the odds of finding a job relative to 15 to 19 year old unemployed individuals. This result highlights the desperate position of youth unemployed and the importance of initiatives, such as the Youth Guarantee, which targets education and training assistance towards younger claimants.

Education was important in securing an exit to employment in both periods, but its importance rose substantially between 2006 and 2011, particularly with respect to third-Level qualifications. Prior to the recession, claimants with third-Level certificates/diplomas and degrees had between 6 and 10 times the odds of finding employment relative to individuals with a Junior Certificate education (lower secondary) or lower, which is quite large. However, by 2011 those with a third-Level certificate/diploma qualification were almost 40 times more likely to find a job relative to those with a Junior Certificate education or lower, while those with a third-Level degree were 126 times more likely to obtain employment. These results suggest that new and replacement labour demand in Ireland post-recession was heavily skewed towards graduates.

In 2006, we observe the well documented scarring effect of previous unemployment durations on unemployment-to-employment labour market transitions. Relative to individuals who were unemployed for less than three months, those with durations of between 6 to 12 months or more than 12 months had around two-thirds (0.66) and one-third (0.37), respectively, the odds of moving into employment by the following quarter. By 2011, however, we find that previous unemployment duration had no discernible impact on exits to employment. One possible explanation for the disappearance of the scarring effect is that the magnitude, and indiscriminate nature, of the incidence of both unemployment and long-term unemployment over the period eradicated the value of durations as a quality signal to employers. Alternatively, the rapid rise in the population of long-term unemployed individuals over the recession with high levels of labour market attachment may have led to a structural behavioural shift in the intensity of job-search among long-term claimants which will also have dampened the scarring effect.

The QNHS contains a number of variables that capture the extent to which an unemployed individual is awaiting a response on a job application and the nature of any job-search activity undertaken in the previous quarter. In 2006, individuals awaiting the results from a public sector competition or an aptitude test had between 1.3 and 1.4 times the odds of exiting to employment by the following quarter relative to unemployed individuals who were not awaiting a response to one of these two job application activities. In 2006, none of the job search variables had any significant impact on exits from unemployment, with the exception of contact with a public employment office: such individuals had 0.8 times the odds of finding a job relative to unemployed individuals that did not use this type of job search activity.^{11,12} Differences in job search activity became somewhat more important in the later period, with individuals who

Table 3 Odds ratios from random effects logit estimations of transitions from unemployment to employment

	2006		2011	
	Odds ratio	Standard error	Odds ratios	Standard error
Male	1.902***	0.195	0.611***	0.069
Age 20-24	5.063***	0.996	9.888***	3.226
Age 25-34	5.138***	1.013	33.055***	11.267
Age 35-44	4.559***	0.936	37.615***	13.202
Age 45-54	5.541***	1.195	44.614***	16.041
Age 55+	3.821***	0.855	26.761***	9.792
Leaving certificate	4.445***	0.622	10.605***	1.726
Post Leaving Cert level (includes apprenticeships)	4.761***	0.913	6.225***	1.099
Third level: non degree	5.909***	1.211	39.703***	9.365
Third level: degree	10.097***	2.059	125.744***	30.770
Irish	1.652***	0.143	2.857***	0.485
Married	1.245*	0.146	2.581***	0.336
Border	0.628***	0.100	0.660*	0.151
Mid-East	0.937	0.179	0.966	0.211
Midlands	0.703	0.155	1.092	0.264
Mid-West	0.501***	0.089	1.009	0.210
South-East	0.762*	0.125	0.728	0.142
South-West	0.705**	0.108	1.228	0.229
West	0.698*	0.133	0.962	0.200
Prev. unemployed for 3–6 months	0.876	0.111	1.217	0.206
Prev. unemployed for 6–12 months	0.658***	0.084	1.145	0.186
Prev. unemployed for 12+ months	0.368***	0.043	0.908	0.125
Prev. sector: Agriculture	1.402	0.576	1.709	0.813
Prev. sector: Construction	0.967	0.165	1.234	0.225
Prev. sector: Wholesale and retail	1.031	0.182	1.435*	0.298
Prev. sector: Transport and storage	0.862	0.255	1.023	0.327
Prev. sector: Accommodation and food service activities	0.825	0.173	1.121	0.283
Prev. sector: Information and communication	0.731	0.239	0.895	0.361
Prev. sector: Financial, insurance and real estate activities	0.797	0.286	0.790	0.358
Prev. sector: Professional, scientific and technical activities	0.694	0.250	0.817	0.268
Prev. sector: Administrative and support service activities	0.990	0.237	1.248	0.372
Prev. sector: Public Administration and defence	0.703	0.315	0.581	0.295
Prev. sector: Education	1.387	0.410	1.536	0.554
Prev. sector: Health	0.901	0.228	1.093	0.326
Prev. sector: Other	1.247	0.312	1.081	0.316
Prev. sector: No sector information	0.866	0.138	1.426*	0.283
Prev. job search: Contacted public employment office	0.819*	0.090	0.656***	0.073
Prev. job search: Contacted private employment agency	1.075	0.114	1.554***	0.176
Prev. job search: Applied directly to employers	0.972	0.116	1.164	0.154
Prev. job search: Asked friends/relatives/trade unions	0.846	0.103	0.643***	0.107
Prev. job search: Studied ads in newspapers	1.009	0.165	0.593***	0.115
Prev. job search: Inserted/answered ads in newspapers	0.991	0.105	0.977	0.108

Table 3 Odds ratios from random effects logit estimations of transitions from unemployment to employment (Continued)

Prev. job search: Looked for permits, licences or financial resources	0.526	0.303	1.274	1.064
Prev. job search: Looked for land, premises or equipment	1.317	0.885	0.987	1.073
Pending job decisions: Waiting for call from public employment office	0.866	0.098	1.021	0.126
Pending job decisions: Waiting for public sector competition results	1.306***	0.136	1.540***	0.190
Pending job decisions: Took a test, interview or exam	1.396***	0.146	1.051	0.131
Pending job decisions: Waiting for results from job application	0.956	0.210	0.874	0.288
Constant	0.052***	0.017	0.000***	0.000
Number of observations	8,299		18,488	
Number of people	6,653		13,776	
Log likelihood	-4820.0517		-9450.0625	
Prob > chi2	0.0000		0.0000	

Note: Dependent variable is transitioned from unemployment to employment. Omitted categories are: age 15–19, Junior Certificate or lower education, Dublin, previously unemployed for less than three months and manufacturing.

Note: ***p < .001; **p < .01; *p < .05.

had interacted with a private employment agency having one and a half times the odds of entering employment by the following quarter relative to those who did not engage with a private recruitment group in 2011. The same result emerged for those waiting the results from a public sector employment competition in 2011. On the other hand, unemployed individuals who had engaged with a public employment office, sought work through family or friends or by searching newspaper advertisements all had lower odds of finding a job within three months relative to unemployed people that did not utilise these job search methods.¹³

With respect to gender, males had almost twice the odds of exiting to employment in 2006 relative to females. However, by 2011, the situation had reversed, with males having around two-thirds (0.61) the odds of moving from unemployment to employment than females. In both time periods examined in this study, Irish nationals that were unemployed were more likely to transition out of unemployment than immigrants, with the scale of the advantage rising in 2011. Neither pre nor post the Great Recession did previous industrial sector employment have any importance in explaining transitions into employment. Married individuals were more likely to successfully exit from unemployment in both periods. Finally, while unemployment-to-employment transition rates were generally lower in most regions relative to Dublin in 2006, apart from the Border region, the remaining regional effects had dissipated by the later period.

Table 4 shows the impact of characteristics on the risk of unemployment in both 2006 and 2011. Again, we report the odds ratios associated with a range of attributes for both time periods. It is important to note that we cannot distinguish the extent to which transitions from employment-to-unemployment were as a consequence of involuntary or voluntary separations. Nevertheless, it is reasonable to assume that the majority of separations were involuntary in nature, particularly in 2011 when the Irish economy was still experiencing substantial difficulties. As was the case with unemployment-to-employment transitions, we observe substantial change in the impacts of age and gender over the period. In 2006, relative to 15 to 19-year-olds, employees aged 20 to 34 had a higher risk of unemployment, as measured by higher odds ratios, while those aged 35 and

Table 4 Odds ratios from random effects logit estimations of transitions from employment to unemployment

	2006		2011	
	Odds ratio	Standard error	Odds ratio	Standard error
Male	1.448***	0.101	3.178***	0.186
Age 20-24	1.861***	0.237	5.388***	0.712
Age 25-34	1.028	0.129	3.466***	0.427
Age 35-44	0.756**	0.102	2.407***	0.305
Age 45-54	0.659***	0.092	1.604***	0.209
Age 55+	0.271***	0.045	0.745**	0.107
Leaving certificate	0.613***	0.055	0.638***	0.050
Post Leaving Cert Level (includes apprenticeships)	0.640***	0.080	1.134	0.102
Third level: non degree	0.589***	0.076	0.377***	0.038
Third level: degree	0.381***	0.044	0.220***	0.021
Irish	1.038	0.060	0.552***	0.044
Married	0.351***	0.031	0.381***	0.025
Border	1.002	0.118	1.202	0.134
Mid-East	0.558***	0.074	1.313***	0.132
Midlands	0.569***	0.091	1.897***	0.228
Mid-West	0.933	0.120	1.901***	0.190
South-East	0.851	0.102	1.773***	0.175
South-West	0.735***	0.079	1.141	0.101
West	0.801	0.110	1.981***	0.199
Prev. sector: Agriculture	0.432***	0.087	0.520***	0.082
Prev. sector: Construction	1.197	0.140	1.061	0.123
Prev. sector: Wholesale and retail	0.904	0.104	0.940	0.087
Prev. sector: Transport and storage	0.894	0.150	0.947	0.125
Prev. sector: Accommodation and food service activities	1.424**	0.200	0.940	0.108
Prev. sector: Information and communication	0.662**	0.134	1.236	0.175
Prev. sector: Financial, insurance and real estate activities	0.671**	0.120	0.807	0.115
Prev. sector: Professional, scientific and technical activities	0.542***	0.099	0.923	0.125
Prev. sector: Administrative and support service activities	1.509**	0.252	1.105	0.158
Prev. sector: Public Administration and defence	0.553***	0.104	0.724**	0.104
Prev. sector: Education	0.885	0.139	0.983	0.120
Prev. sector: Health	0.975	0.125	1.117	0.110
Prev. sector: Other	1.123	0.179	0.895	0.117
Prev. sector: No sector information	0.715	0.411	0.415	0.291
Prev. Permanent contract	1.100	0.085	1.101	0.068
Prev. Full-time	0.882*	0.057	0.683***	0.035
Prev. Trade union member	0.812**	0.067	0.845**	0.058
Prev. Firm size 11-19	1.088	0.112	1.007	0.089
Prev. Firm size 20-49	0.917	0.092	0.970	0.079
Prev. Firm size 50-99	1.103	0.126	0.865	0.083
Prev. Firm size 100-500	1.018	0.102	0.915	0.077
Prev. Firm size 500+	1.081	0.151	0.823*	0.083
Constant	0.001***	0.000	0.004***	0.001

Table 4 Odds ratios from random effects logit estimations of transitions from employment to unemployment (Continued)

Number of observations	178,325	104,879
Number of people	91,815	57,862
Log likelihood	-14706.688	-18397.634
Prob > chi2	0.0000	0.0000

Note: Dependent variable is transitioned from employment to unemployment. Omitted categories are: age 15–19, Junior Certificate or lower education, Dublin, manufacturing and firm size 1–10 employees.

Note: ***p < .001; **p < .01; *p < .05.

over had lower odds of becoming unemployed - with the risk falling with age. In 2011, employees aged between 20 and 54 had between 1.6 and five times the odds of becoming unemployed relative to the younger cohort. Only workers aged 55 and over had a lower risk of unemployment in 2011 compared to those aged 15 to 19. A potential explanation for this result is that firms became more sensitive to the price of labour post-recession, implying that the youngest workers, who also tend to be the least expensive, had the lowest risk of a redundancy arising from any cost rationalisation process. Nevertheless, it must be borne in mind that younger workers, due to their relative lack of experience are also likely to be less productive, suggesting that any substitution for more experienced labour is likely to be restricted to relatively unskilled occupations. Thus, while 15 to 19-year-olds are at the greatest risk of sustained unemployment, they appear to have a relatively low risk of job loss.

With respect to education, we observe a general linear relationship in both periods, with the risk of unemployment declining with higher levels of educational attainment. Relative to 2006, third-level qualifications are of greater importance in deterring unemployment in 2011, which is again consistent with a change in the composition of labour demand in favor of educated workers. Furthermore, by 2011, the Post Leaving Cert (PLC) level qualification, which is typically vocational in nature, no longer afforded workers any advantage relative to those with a Junior Certificate or lower levels of education. In 2006, males had approximately 1.5 times the odds of transitioning to unemployment relative to females. However, by 2011, males had over three times the odds of transitioning into unemployment compared to females. In the earlier period, Irish workers and immigrants had a similar risk of becoming unemployed, but by 2011, Irish natives had approximately half the odds of moving into unemployment relative to immigrants.

In 2006, relative to manufacturing, the risk of moving into unemployment was higher for workers in accommodation and food services as well as administrative and support services; and it was lower for those in the agriculture, public administration, professional, financial, and information and communication sectors. By 2011, only workers in the public administration sector had a lower likelihood of becoming unemployed relative to those working in manufacturing. In both 2006 and 2011, full-time workers, and those who were members of trade unions, had lower odds of unemployment, with the impact relatively constant over time. Finally, with respect to region relative to Dublin, the risk of unemployment was lower in a number of areas in 2006 (e.g., Mid-East and Midlands). However, by 2011, employees located in areas other than Dublin had higher odds of becoming unemployed. This result suggests a greater concentration of economic activity in the capital post-recession.

The random effects logit models gave us a sense of the extent to which the value of certain attributes changed pre and post recession with respect to transitioning from, or falling into, unemployment. It is not clear, however, the extent to which variations in the transition rate over time are being driven by (i) changes in the value of observable characteristics, (ii) changes in the composition of the populations in employment and unemployment, or (iii) unobservables that may relate to either individuals or changes in the external environment. In order gain further insights, we estimated the following non-linear Oaxaca decomposition, on the pooled years of data, of the change in transition rates from 2006 to 2011:

$$\overline{T}_{11} - \overline{T}_{06} = (\overline{X}_{11} - \overline{X}_{06})\hat{\beta}_{111} + (\hat{\beta}_{111} - \hat{\beta}_{106})\overline{X}_{06} + (\hat{\alpha}_{11} - \hat{\alpha}_{06}) \quad (2)$$

The term on the left-hand-side, $\overline{T}_{11} - \overline{T}_{06}$, captures the change in the transition rate between 2006 and 2011. The first term on the right-hand-side is known as the endowment effect: this effect captures the proportion of the change in the transition rate between 2006 and 2011 that is due to differences in the composition of the employed and unemployed populations in both time periods. The second term captures the coefficient effect, which is a measure of the extent to which the 2006 and 2011 employed and unemployed populations holding similar characteristics are treated differently in the labour market. The final term in equation (2) denotes the factors that we cannot control for in our model that will impact on a person's transition rate from employment to unemployment and unemployment to employment. These factors include, for example, social class, previous employment tenure, etc., (i.e., unobservables). The results from this decomposition analysis are shown in Table 5.

The first observation to note is that endowment effects explained very little of the overall change in either transition rate examined, suggesting that changes in the composition of the respective populations over the period mattered little. While the structure of employment will tend to change slowly over time, the lack of an endowment effect on the transition from unemployment-to-employment is somewhat surprising given the well documented rapid growth in the proportions of unemployed claimants with higher levels of education since the onset of the recession.

Dealing firstly with the transitions from unemployment-to-employment, the vast majority of the change in the transition rate can be attributed to changes in the constant coefficient, which will incorporate largely unobserved external factors. Nevertheless, some changes in the value of certain observable characteristics had important impacts

Table 5 Decomposition results for changes in transition rates

	Employment to unemployment	Unemployment to employment
Overall difference	0.0311	-0.0921
Endowment effect	-0.0103	0.0277
Coefficient effect:	0.0325	-0.1052
Of which:		
Observables	0.0105	0.0358
Constant	0.0220	-0.1410
Interaction effect	0.0089	-0.0146

Note: The decomposition is one for binary dependent variables proposed by Yun (2004). The analysis uses pooled data for 2006 and 2011 and 2006 is the base year in both decompositions.

on the changing rate of transition. The decline in the impact of being male is the most important single coefficient effect, which resulted in a 6 percentage point reduction in the transition rate from unemployment-to-employment between 2006 and 2011.¹⁴ It is not clear what the gender coefficient effect is picking up, given that the models account for claimants' previous industrial sector. However, it is not necessarily the case that the control for previous sector of employment will reflect the degree to which the human capital of claimants meets the requirements of current labour demand. Given this, the results would suggest that between 2006 and 2011, the demand for labour dramatically switched towards industries and occupations where females were more heavily represented, with little evidence that claimants previously employed in particular sectors of employment heavily affected by the economic downturn were at any particular disadvantage in accessing emerging opportunities. In support of this view, McGinnity *et al.* (2014) show that the Health and Social Work sector, which is heavily female dominated, exhibited the largest absolute level of sectoral employment growth¹⁵ in Ireland over the 2007 to 2012 period.¹⁶

The rise in the coefficients for older unemployed individuals, which presumably proxies an increase in the value of labour market experience, exerted a positive influence on the change in the transition rate. The very substantial reduction in the scarring impact of long-term unemployment durations also tended to increase transition rates over the period. The elimination of the decline in the relative return to the vocationally orientated PLC qualification had a depressing affect on the transition rate, and it is again likely that the fall in the return to this qualification is correlated with the lack of demand in vocationally orientated sectors such as construction, manufacturing and wholesale and retail.

With respect to the rise in the rate of transitions to unemployment,¹⁷ the increase in the relative exposure of males and older workers were the largest single coefficient effects. In this instance, the age effect may well reflect a rise in the price sensitivity of employers to labour costs post recession.

5 Policy implications

From a policy perspective, the results emphasize the importance of higher level credentials in terms of preventing transitions to unemployment and exiting unemployment. The research is suggestive of a rise in demand for graduate labour over time, which would be supportive of a further expansion of higher-education provision. In terms of activation policy, there has been a substantial focus on enabling unemployed claimants to gain third-level qualifications through the Springboard initiative, which currently offers 6,400 free higher education places to jobseekers leading to qualifications at certificate, degree and post-graduate level.

The study also identifies that labour demand is evolving in an uneven pattern with respect to both male and female dominated industries. Over the period in question, employment growth appears to have been particularly skewed towards female dominated sectors, which has resulted in a substantial lowering in the likelihood that males will exit the Live Register¹⁸ in any three-month period. Although no significant sectoral effects were obvious at the one-digit industry-level, the results point to the possible existence of structural unemployment within the male claimant population, with the risk likely to be high for

persons previously employed in particular construction occupations. Activation policy has, to date, failed to identify or address the degree to which particular occupations have been disproportionately disadvantaged in terms of re-entering employment, and the results suggest that more action is required in this respect.

With respect to labour market activation and young people, the research highlights the drastic situation faced by unemployed youths. Thus, the findings emphasise the importance of implementing the Youth Guarantee. Nevertheless, the scale of the unemployment problem in Ireland and the associated strain on resources implies that a much more scaled down version of the Guarantee will be implemented to what was initially envisaged. Specifically, claimants under the age of 18 who leave school without formal qualifications will be offered a “second chance” education place outside the school system; however, this aspect of the Guarantee will not be fully operational until the end of 2015. Claimants aged between 18 and 24 will be categorised as high and low risk on the basis of the country’s profiling system. Under this system, each new claimant is asked a series of questions on making a claim for unemployment benefit, and a probability of exit to the labour market (PEX) score is calculated. Individuals with a low PEX will be deemed high risk. High risk claimants will only receive an intervention after four months on the Live Register, while the waiting time for medium- to high-risk claimants will be nine months. It is likely that such relatively low intensity interventions will have limited impact on the transition rates of young claimants to unemployment.

Finally, the apparent reduction of the scarring effect associated with long-term unemployment suggests that activation policy need not be so heavily focused around durations, but instead should attempt to ensure that the claimants are equipped with the right skills and provided with the appropriate advice and training that would enable them to re-enter employment irrespective of their length of time spent in unemployment.

6 Conclusions

This paper assesses the impact of the Irish recession on transitions to and from unemployment.¹⁹ We find, not surprisingly, that the rate of transition from unemployment to employment declined over the period, while the transition rate from employment to unemployment increased.

In terms of the factors contributing to the likelihood of a transition taking place, some important changes appear to have occurred over the 2006 to 2011 period. The results show that young people are much less likely to exit unemployment while simultaneously having a lower risk of becoming unemployed. Education has become an increasingly important factor post recession in both securing an exit from unemployment and preventing transitions to unemployment, with the impact most pronounced for the transition from unemployment-to-employment. We find that the scarring effect of long-term unemployment has fallen substantially post-recession, suggesting that either (i) the rapid and indiscriminate rise in the incidence of long-term durations has eradicated the signalling value of the attribute, or (ii) structural shifts in the population of long-term claimants has led to a rise in the job-search intensity causing a decline in the relationship between unemployment durations and exits to employment.

Despite the rapid changes in the structure of the unemployed population, results from a decomposition analysis show that compositional changes are largely unimportant in explaining the fall in the transition rate from unemployment to employment between

2006 and 2011. While most of the change is, not surprisingly, attributable to unobserved changes in the external environment, variations in the value of certain attributes also explained some of the movement in transition rates. In particular, we find that the fall in the return of being male over time contributed to a decline in the transition rate from unemployment to employment, with the gender effect indicating that labour demand was more heavily skewed towards female dominated sectors over the 2006 to 2011 period. The decline in the scarring effect of long-term unemployment and the relative rise in the return to labour market experience, proxied by age, exerted a positive influence on the transition rate to employment over the period.

From a policy perspective, the study highlights the importance of initiatives that focus on equipping claimants with educational credentials: while existing programmes that provide routes into third-level schooling, such as Springboard, appear fit for purpose, aspects of the Youth Guarantee appear under-resourced relative to the scale of the problem. Finally, the apparent reduction of the scarring effect associated with long-term unemployment suggests that activation policy need not be so heavily focused around targeting durations but, instead, should concentrate on ensuring claimants are properly equipped with the correct information and skills that will facilitate re-entry to employment. As Ireland continues to grow, it is crucial that the country's Public Employment Services are underpinned by accurate labour market intelligence that identifies areas of emerging employment growth and that existing training and education programmes are sufficiently flexible to respond the changing needs of the economy.

Endnotes

¹This is a post-secondary, mainly vocational, qualification.

²Eurostat, March 2013: http://appsso.eurostat.ec.europa.eu/nui/show.do?dataset=edat_lfse_20&lang=en.

³Our analysis does not include inactivity either in the transition to or from unemployment. While there was a marked increase in the rate of inactivity over the Great Recession, the focus of this paper is solely on movements from and into unemployment given the dramatic increase in unemployment following the downturn.

⁴The QNHS is carried out by the Central Statistics Office (CSO), Ireland's national statistical collection organisation.

⁵The CSO introduced a new sample in Quarter 4 2012 as a result of the 2011 Census of Population. This new sample was introduced on an incremental basis across each quarter from Quarter 4 2012 to Quarter 4 2013. Thus, the new sample has only been fully effective since Quarter 4 2013.

⁶Information provided by the CSO.

⁷The two time-points considered in the analysis broadly approximate peak to trough activity in the labour market. Of course, there could also be rapid changes within the two time periods considered here. However, this analysis is concerned with how the nature of transitions changed from peak to trough activity in Ireland's labour market.

⁸Ward-Warmedinger and Macchiarelli (2014) examined labour market transition patterns across a range of EU regions (Central Eastern, Nordics, Continental and Mediterranean) over the period 2004 to 2008: overall (see Table 2) and broken down by gender, age and educational attainment (see Table 4). Their results are not directly comparable with those derived in this paper due to variations in the time periods examined and the use of different

Table 6 Oaxaca individual coefficient effects for the change in the unemployment to employment transition rate between 2006 and 2011

	Coefficient	Standard error
Male	−0.060***	(0.007)
Age 20-24	−0.004	(0.004)
Age 25-34	0.012*	(0.007)
Age 35-44	0.014**	(0.006)
Age 45-54	0.011**	(0.005)
Age 55+	0.008**	(0.003)
Leaving certificate	−0.004	(0.004)
Post leaving cert level (includes apprenticeships)	−0.008**	(0.003)
Third level: non degree	0.004	(0.002)
Third level: degree	0.006**	(0.003)
Irish	−0.019	(0.012)
Married	0.011**	(0.006)
Border	0.002	(0.002)
Mid-East	0.001	(0.002)
Midlands	0.002	(0.002)
Mid-West	0.008***	(0.003)
South-East	0.001	(0.003)
South-West	0.008**	(0.003)
West	0.003	(0.003)
Prev. unemployed for 3–6 months	0.003*	(0.002)
Prev. unemployed for 6–12 months	0.006**	(0.003)
Prev. unemployed for 12+ months	0.051***	(0.009)
Prev. sector: Agriculture	0.000	(0.001)
Prev. sector: Construction	0.001	(0.005)
Prev. sector: Wholesale and retail	0.001	(0.003)
Prev. sector: Transport and storage	0.001	(0.001)
Prev. sector: Accommodation and food service activities	0.002	(0.002)
Prev. sector: Information and communication	0.000	(0.001)
Prev. sector: Financial, insurance and real estate activities	0.000	(0.001)
Prev. sector: Professional, scientific and technical activities	0.000	(0.001)
Prev. sector: Administrative and support service activities	0.000	(0.001)
Prev. sector: Public Administration and defence	0.000	(0.001)
Prev. sector: Education	−0.001	(0.001)
Prev. sector: Health	0.000	(0.001)
Prev. sector: Other	−0.001	(0.001)
Prev. sector: No sector information	0.005	(0.004)
Prev. job search: Contacted public employment office	−0.006	(0.008)
Prev. job search: Contacted private employment agency	0.005	(0.004)
Prev. job search: Applied directly to employers	0.014	(0.013)
Prev. job search: Asked friends/relatives/trade unions	−0.008	(0.016)
Prev. job search: Studied ads in newspapers	−0.030	(0.021)
Prev. job search: Inserted/answered ads in newspapers	−0.002	(0.004)
Pending job decisions: Waiting for call from public employment office	0.004	(0.004)

Table 6 Oaxaca individual coefficient effects for the change in the unemployment to employment transition rate between 2006 and 2011 (Continued)

Pending job decisions: Waiting for results from job application	0.000	(0.001)
Prev. job search: Looked for permits, licences or financial resources	0.000	(0.000)
Prev. job search: Looked for land, premises or equipment	0.000	(0.000)
Pending job decisions: Waiting for public sector competition results	0.002	(0.003)
Pending job decisions: Took a test, interview or exam	−0.007**	(0.003)
Constant	−0.141***	(0.040)

Note: ***p < .001; **p < .01; *p < .05.

employment and unemployment definitional measures (ILO versus subjective measures). Nevertheless, a loose comparison of the results suggests that the employment-to-unemployment transition rate in Ireland in 2006, which would be similar to the pre-recessionary period of 2004–2008 examined by Ward-Warmedinger and Macchiarelli (2014), was lower than the EU regional rates identified by Ward-Warmedinger and Macchiarelli (2014): the average rate for the regions examined was 2.7 per cent, which compares with 1.8 per cent for Ireland. For the same time periods (2006 compared to 2004–2008), the unemployment-to-employment transition rate in Ireland was higher (37.3 per cent) for all the EU areas examined by Ward-Warmedinger and Macchiarelli (2014), apart from the Nordic region's rate (39.7 per cent). The age and educational attainment labour market transition rate results derived by Ward-Warmedinger and Macchiarelli (2014) are again not directly comparable with those derived in this paper due to the reasons outlined previously, and also because different age and educational attainment definitions were employed in both studies.

⁹The data in our analysis spans a relatively short time period; individuals can only be included in the panels for a maximum of five quarters. As such, alternative techniques such as duration models, which generally require longer panels of data with lower attrition rates, are not suitable for this analysis.

¹⁰Although our models allow for individuals to be in the analysis for up to five quarters, on average, individuals are present in the unemployment to employment transition models in 2006 and 2011 for 1.2 and 1.3 quarters respectively, and for 1.9 and 1.8 quarters in the employment to unemployment transitions models in 2006 and 2011 respectively. This minimises the possibility that our results are affected by individuals making multiple transitions over the five-quarter periods; in fact, excluding individuals who make multiple transitions has no material impact on the model results. These results are available on request from the authors.

¹¹This finding is consistent with the work of McGuinness et al. (2011) who report that the activation system in force at the time (2006), the National Employment Action Plan (NEAP), actively reduced exits from the Live Register to employment.

¹²In addition, to serve as a proxy for job-search intensity, a variable was created that sums the number of different job search activities undertaken by an individual. However, the estimated odds ratio was close to one in both models and insignificant.

¹³With respect to the public employment office, while the NEAP was in the process of being replaced in 2011, the scale of the unemployment problem at that time meant that existing resources were severely stretched in their ability to deal with the unemployment crisis that existed in the country at that time.

Table 7 Oaxaca individual coefficient effects for the change in the employment to unemployment transition rate between 2006 and 2011

	Coefficient	Standard error
Male	0.007***	(0.001)
Age 20-24	0.001***	(0.000)
Age 25-34	0.005***	(0.001)
Age 35-44	0.004***	(0.001)
Age 45-54	0.003***	(0.001)
Age 55+	0.002***	(0.001)
Leaving Certificate	0.000	(0.001)
Post Leaving Cert Level (includes apprenticeships)	0.001***	(0.000)
Third level: non degree	-0.001***	(0.000)
Third level: degree	-0.003***	(0.001)
Irish	-0.013***	(0.002)
Married	0.000	(0.001)
Border	0.000	(0.000)
Mid-East	0.002***	(0.000)
Midlands	0.001***	(0.000)
Mid-West	0.001***	(0.000)
South-East	0.002***	(0.000)
South-West	0.002***	(0.000)
West	0.002***	(0.000)
Prev. sector: Agriculture	0.000	(0.000)
Prev. sector: Construction	0.000	(0.000)
Prev. sector: Wholesale and retail	0.000	(0.000)
Prev. sector: Transport and storage	0.000	(0.000)
Prev. sector: Accommodation and food service activities	-0.001***	(0.000)
Prev. sector: Information and communication	0.000***	(0.000)
Prev. sector: Financial, insurance and real estate activities	0.000	(0.000)
Prev. sector: Professional, scientific and technical activities	0.000*	(0.000)
Prev. sector: Administrative and support service activities	0.000	(0.000)
Prev. sector: Public Administration and defence	0.000	(0.000)
Prev. sector: Education	0.000	(0.000)
Prev. sector: Health	0.000	(0.000)
Prev. sector: Other	0.000	(0.000)
Prev. sector: No sector information	0.000	(0.000)
Prev. Permanent contract	-0.001	(0.001)
Prev. Full-time	-0.003***	(0.001)
Prev. Trade union member	0.000	(0.001)
Prev. Firm size 11-19	0.000	(0.000)
Prev. Firm size 20-49	0.000	(0.000)
Prev. Firm size 50-99	0.000	(0.000)
Prev. Firm size 100-500	0.000	(0.000)
Prev. Firm size 500+	0.000	(0.000)
Constant	0.022***	(0.004)

Note: ***p < .001; **p < .01; *p < .05.

¹⁴See Table 6.

¹⁵Sectoral employment increased by 23,585 (10.6 per cent).

¹⁶Over the period, total employment fell by 303,091 (14.1 per cent).

¹⁷See Table 7.

¹⁸Ireland's national register of all individuals in receipt of unemployment benefits.

¹⁹For an assessment of the income distributional impact of the recession in Ireland, see Callan et al. (2014).

Competing interests

The IZA Journal of European Labor Studies is committed to the IZA Guiding Principles of Research Integrity. The authors declare that they have observed these principles.

Acknowledgements

We would like to thank the Central Statistics Office (CSO) for making the Quarterly National Household Survey (QNHS) data available to us and, in particular, Brian Ring (CSO) for his assistance with data queries. The authors would like to thank the anonymous referee.

Responsible editor: Martin Kahanec

Author details

¹The Economic and Social Research Institute, Whitaker Square, Sir John Rogerson's Quay, Dublin 2, Ireland. ²The Department of Economics, Trinity College Dublin, Dublin, Ireland. ³National Institute for Labour Studies (NILS), Flinders University, Adelaide, Australia.

Received: 15 October 2014 Accepted: 28 February 2015

Published online: 29 April 2015

References

- Barrett A, Kelly E (2012) The impact of Ireland's recession on the labour market outcomes of its immigrants. *Eur J Popul* 28(1):99–111
- Callan T, Nolan B, Keane C, Savage M (2014) Walsh JR (2014) Crisis, response and distributional impact: the case of Ireland. *IZA J Eur Labor Stud* 3:9
- Conefrey T (2011) Unemployment and Labour Force Participation During the Recession, Economic Letter Series from Central Bank of Ireland. 4
- Kelly E, McGuinness S (2014) The impact of the recession on the structure and labour market success of NEET youth in Ireland. *Intereconomics* 48(4):196–235
- Kelly E, McGuinness S, O'Connell PJ (2012) Submission to the Joint Committee on Jobs, Social Protection and Education on unemployment and youth unemployment
- Kelly E, McGuinness S, O'Connell PJ, Haugh D (2014) Pandiella González A (2014) Transitions in and out of unemployment among young people in the Irish recession. *Comp Econ Stud* 56:616–634
- McGinnity F, Russell H, Watson D, Kingston G, Kelly E (2014) Winners and losers? The equality impact of the great recession in Ireland. The Equality Authority and the Economic and Social Research Institute, Dublin
- McGuinness S, O'Connell PJ, Kelly E, Walsh JR (2011) Activation in Ireland: An Evaluation of the National Employment Action Plan. ESRI Research Series No. 20. Economic and Social Research Institute, Dublin
- Ward-Warmedinger M, Macchiarelli C (2014) Transitions in labour market status in EU labour markets. *IZA J Eur Labor Stud* 3:17
- Yun MS (2004) Decomposing differences in the first moment. *Econ Lett* 82:275–280

Submit your manuscript to a SpringerOpen[®] journal and benefit from:

- Convenient online submission
- Rigorous peer review
- Immediate publication on acceptance
- Open access: articles freely available online
- High visibility within the field
- Retaining the copyright to your article

Submit your next manuscript at ► springeropen.com